SYSTEMATIC REVIEW



Efficacy of Multi-Component Exercise-Based Injury Prevention Programs on Injury Risk Among Footballers of All Age Groups: A Systematic Review and Meta-analysis

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Abstract

Background Playing football is associated with a high risk of injury. Injury prevention is a priority as injuries not only negatively impact health but also potentially performance. Various multi-component exercise-based injury prevention programs for football players have been examined in studies.

Objective We aimed to investigate the efficacy of multi-component exercise-based injury prevention programs among footballers of all age groups in comparison to a control group.

Methods We conducted a systematic review and meta-analysis of randomized and cluster-randomized controlled trials. CINAHL, Cochrane, PubMed, Scopus, and Web of Science databases were searched from inception to June 2022. The following inclusion criteria were used for studies to determine their eligibility: they (1) include football (soccer) players; (2) investigate the preventive effect of multi-component exercise-based injury prevention programs in football; (3) contain original data from a randomized or cluster-randomized trial; and (4) investigate football injuries as the outcome. The risk of bias and quality of evidence were assessed using the Cochrane Risk of Bias Tool and the Grading of Recommendations Assessment, Development, and Evaluation (GRADE), respectively. The outcome measures were the risk ratio (RR) between the intervention and the control group for the overall number of injuries and body region-specific, contact, and non-contact injuries sustained during the study period in training and match play.

Results Fifteen randomized and cluster-randomized controlled trials with 22,177 players, 5080 injuries, and 1,587,327 exposure hours fulfilled the inclusion criteria and reported the required outcome measures. The point estimate (RR) for the overall number of injuries was 0.71 (95% confidence interval [CI] 0.59–0.85; 95% prediction interval [PI] 0.38–1.32) with very low-quality evidence. The point estimate (RR) for lower limb injuries was 0.82 (95% CI 0.71–0.94; 95% PI 0.58–1.15) with moderate-quality evidence; for hip/groin injuries, the RR was 0.56 (95% CI 0.30–1.05; 95% PI 0.00–102.92) with low-quality evidence; for knee injuries, the RR was 0.69 (95% CI 0.52–0.90; 95% PI 0.31–1.50) with low-quality evidence; for ankle injuries, the RR was 0.73 (95% CI 0.55–0.96; 95% PI 0.36–1.46) with moderate-quality evidence; and for hamstring injuries, the RR was 0.83 (95% CI 0.50–1.37) with low-quality evidence. The point estimate (RR) for contact injuries was 0.70 (95% CI 0.56–0.88; 95% PI 0.40–1.24) with moderate-quality evidence, while for non-contact injuries, the RR was 0.78 (95% CI 0.55–1.10; 95% PI 0.25–2.47) with low-quality evidence.

Conclusions This systematic review and meta-analysis indicated that the treatment effect associated with the use of multicomponent exercise-based injury prevention programs in football is uncertain and inconclusive. In addition, the majority of the results are based on low-quality evidence. Therefore, future high-quality trials are needed to provide more reliable evidence.

Clinical Trial Registration PROSPERO CRD42020221772.

Extended author information available on the last page of the article

Key Points

The present meta-analysis is the first to use prediction intervals in the interpretation of results derived from trials assessing the efficacy of multi-component exercisebased injury prevention programs among footballers of all age groups.

This study revealed that the evidence for meaningful effects of exercise-based injury prevention programs remains inconclusive at best.

The quality of evidence is a major issue in existing studies; therefore, these findings call for future high-quality trials to provide more reliable evidence.

1 Background

The overall injury incidence in professional male football players is between 5.9 [1] and 9.6 [2] injuries/1000 football hours. In amateur and veteran football, reported incidences are even higher and reach 9.6 [2] to 12.5 [3] and 12.4 [4] injuries/1000 football hours, respectively. There are hardly any data regarding players under the age of 11 years [5]. A professional football team with 25 players has approximately 50 injuries per season [6], and youth elite teams about 30 [7]. Many efforts have been made in recent years to reduce these numbers. Various injury prevention programs for football players of both sexes and various age groups have been established. Some of them target specific injuries, for example, Prevent injury and Enhance Performance [8] and HarmoKnee [9], target knee injuries. Others take a more general approach, trying to prevent non-contact lower extremity injuries in general for example, FIFA[®] 11 [10], $FIFA^{(0)}$ 11 + [11], and the Neuromuscular training program [12]. 11 + Kids [13] aims to prevent football injuries by increasing children's fundamental and sport-specific motor skills.

Previous systematic reviews and meta-analyses have evaluated the efficacy of either specific programs (e.g., FIFA 11 and 11+) [14, 15] or the effect of various programs on specific injuries (e.g., non-contact injuries) [16]. However, recognizing the differences between programs regarding the content, the different age groups targeted, and the different results reported compared to each other, a comprehensive meta-analysis of pooled results across the studies will produce a more comprehensive result. To date, no metaanalysis is available that has evaluated the efficacy of all multi-component exercise-based injury prevention programs in reducing the overall number of injuries as well as body region-specific injuries, and considering footballers of all age groups (children, youth, senior, and veteran). Additionally, contact-related injuries represent 50% of overall injuries in professional football [17]. Previous research has not investigated the impact of the programs on preventing these injuries. Providing information about the age-specific efficacy and estimating the potential of these programs on contact-related injuries may guide future evidence-based directions regarding the implementation and development of new interventions. Finally, providing only confidence intervals (CIs) might not be the best way forward. A recent meta-analysis examined the effect of the Nordic hamstring exercise [18]. The authors strongly recommended providing the prediction intervals (PIs) in addition to CIs. This is in line with authors promoting the use of PIs in the interpretation of results from a random-effects meta-analysis of trials assessing treatment effects [19]. Therefore, and for the first time, this meta-analysis reports the PIs in addition to the CIs. The aim of this meta-analysis was to investigate the efficacy of multi-component exercise-based injury prevention programs in reducing injuries of different types among footballers of all age groups.

2 Methods

2.1 Protocol and Registration

We report this systematic review in accordance with the guidelines of the Preferred Reporting Items for Systematic reviews and Meta-Analyses (PRISMA) [20]. The study was registered at PROSPERO (ID: CRD42020221772).

2.2 Study Eligibility Criteria

In the present study, we included all controlled, multi-component exercise-based injury prevention programs containing at least two or more exercises. Players of the intervention group performed these programs during their training sessions in addition to their usual training and were compared to a control group. Criteria for study inclusion were: (1) include football (soccer) players; (2) investigate the preventive effect of multi-component exercise-based injury prevention programs in football; (3) contain original data from a randomized or cluster-randomized trial; and (4) investigate football injuries as the outcome. Studies were excluded from the meta-analysis if they were: (1) studies with a single exercise intervention; (2) studies with a primary target on performance or other physical measurements than injuries; (3) studies using protective equipment (e.g., bracing) as part of the intervention; and (4) studies published in a language other than English.

2.3 Sources and Study Selection

Possible studies were identified using a systematic search process. First, we searched the following databases CINAHL, Cochrane, PubMed, Scopus, and Web of Science from the earliest record to June 2022, with the following search strategy: (injury prevention OR warm-up program OR neuromuscular program OR f-marc OR 11 +) AND (football OR soccer). The reference lists of the studies recovered were hand searched to identify potentially eligible studies missed by electronic searches. Two reviewers independently (AB, DK) performed the selection of studies based on the title and abstract provided by the bibliographic databases. The fulltext evaluation followed on those selected studies from the first selection step. A third reviewer (RO) was responsible for resolving any discrepancies in the selection process.

2.4 Data Extraction and Administration

For each eligible study, four reviewers (RM, AB, DK, AL) extracted data independently using a standardized data extraction form [14]. One section was added (type of injuries: contact or non-contact) to the extraction form for an additional analysis that we performed regarding the effect on contact versus non-contact injuries. We extracted data on the studies' basic information, design, participants, intervention characteristics, and outcome measures. Thereafter, the reviewers compared the extracted data for consistency. Reviewers resolved discrepancies by discussion and, when necessary, a fifth party (RO) was involved. Final decisions were made based on a majority vote. Primary outcome results from individual studies were extracted and collated in Excel 365 (Microsoft Corporation, Redmond, WA, USA).

2.5 Quality Assessment

The risk of bias was assessed for each included trial according to the recommendations outlined in the Cochrane Handbook for Systematic Reviews of Interventions [21]. The following items were considered: allocation sequence generation, concealment of allocation, blinding of outcome assessment, incomplete outcome data, selective outcome reporting, and other sources of bias. As it is impossible to blind the participants to the intervention, we removed the item "blinding of participants and investigators". Each bias domain was judged as at low or high risk of bias according to its possible effect on the results of the study. When the possible effect was unknown or insufficient detail was reported, we judged it as unclear. The risk of bias was examined independently by two reviewers (RO, BSH). Discrepancies were resolved by consensus. The overall quality of evidence was assessed using the Grading of Recommendations Assessment, Development, and Evaluation (GRADE). This method assesses the strength of evidence derived from systematic reviews [22]. In the GRADE system, randomized controlled trials (RCTs) begin as high-quality evidence [23]. Subsequently, the evidence is downgraded by one level for each of the following domains considered: (1) risk of bias (downgraded by one level if the trials scored an overall high risk of bias on the Cochrane Collaboration Risk of Bias Tool); (2) inconsistency (downgraded by one level if statistical heterogeneity between studies was $I^2 > 50\%$; (3) indirectness (downgraded by one level if the meta-analysis included participants with heterogeneous characteristics with regard to sex, age, and level of sport); (4) imprecision (downgraded by one level if the upper and lower CIs had a > 0.5 difference); and (5) publication bias (assessed with a visual inspection of a funnel plot and two-tailed Egger's test if more than ten studies were included in the meta-analysis). Evidence obtained was categorized into four levels of evidence quality: high, moderate, low, and very low [24] (Table 1).

2.6 Outcome Measures

The primary outcome was the risk ratio (RR) for the overall number of injuries. Body region-specific injury RRs for the lower limb, hamstring, hip/groin, knee, and ankle were secondary outcomes. Additionally, the overall number and the region-specific injury RRs were assessed for a non-contact versus contact induced cause. All injuries occurring in official training and match play during the respective study period were included.

2.7 Synthesis of Results

If studies did not report RR estimates, we converted them to RRs as far as possible [25, 26]. Out of the 15 included studies, six studies did not perform cluster adjustments. They also did not provide information on the intra-cluster correlation coefficient or other data that would allow for calculating the design effect or inflation factor (as recommended by the Cochrane Handbook for Systematic Review of Interventions) [27]. Hence, we performed a cluster adjustment by increasing variance by 30% for effect estimates of studies with no adjustment for the cluster effect [28]. We performed a meta-analysis of RRs and their 95% CIs using the DerSimonian and Laird random-effects method [29]. A randomeffects meta-analysis assumes that the true treatment effect varies among studies. The DerSimonian and Laird method does not make any assumptions about the distribution of the random effects [30]. In addition to the presentation of overall effect estimates and 95% CIs, we also calculated 95% PIs.

Meta-analyses	Number of RCT's	Risk of bias ^a	Inconsistency ^b	Indirectness ^c	Imprecision ^d	Publication bias ^e	Effect RR (95% CI)	GRADE quality
Overall injuries	11	0	0	θ	\oplus	0	0.71	$\Theta \Theta O O O$
All studies							(0.59 to 0.85)	Very low
Lower limb injuries	9	0	θ	θ	Φ	-	0.82	$\Theta \Theta \Theta \odot$
All studies							(0.71 to 0.94)	Moderate
Hip/groin injuries	3	0	\oplus	θ	0	-	0.56	$\Theta \Theta \odot \odot$
All studies							(0.30 to 1.05)	Low
Hamstring injuries	2	0	θ	θ	0	-	0.83	$\Theta \Theta \odot \odot$
All studies							(0.50 to 1.37)	Low
Knee injuries	11	0	0	\oplus	\oplus	\oplus	0.69	$\Theta \Theta \Theta \odot \odot$
All studies							(0.52 to 0.90)	Low
Ankle injuries	7	0	\oplus	θ	\oplus	-	0.73	$\Theta \Theta \Theta \odot$
All studies							(0.55 to 0.96)	Moderate
Contact injuries	6	0	θ	θ	Φ	-	0.68	$\Theta \Theta \Theta \odot$
All studies							(0.55 to 0.84)	Moderate
Non-Contact injuries	6	0	0	θ	\oplus	-	0.73	$\Theta \Theta \odot \odot$
All studies							(0.53 to 1.01)	Low

Table 1 Grades of recommendation, assessment, development and evaluation (GRADE) quality of evidence

^a Downgraded if the trials scored an overall high risk of bias on the Cochrane Collaboration Risk of Bias Tool.

^b Downgraded if statistical heterogeneity between studies was (I² > 50%).

^c Downgraded by one level if meta-analysis included participants with heterogeneous characteristics.

^d Downgraded if the upper and lower CIs had >0.5 difference.

^e Assessed with visual inspection of the funnel plot and two-tailed Egger test (if >10 studies were included in the meta-analysis).

CI confidence interval, RCT randomized controlled trial, RR risk ratio

They enable the examination of treatment effects within an individual study setting, as this can differ from the average effect [19]. Heterogeneity was assessed using I^2 , τ^2 , and Q value (χ^2 test for heterogeneity). We interpreted I^2 values according to guidelines by Higgins and Green, a low heterogeneity for I^2 values between 25 and 50%, a moderate f heterogeneity or 50–75%, and a high heterogeneity for $\geq 75\%$ [27]. A small study effect was investigated using Egger's test for a meta-analysis with ten or more studies [31]. Statistical analysis was carried out using STATA 17 BE (Stata Corporation, College Station, TX, USA).

3 Results

3.1 Literature Identification

The initial database search identified 7954 studies. Following the removal of duplicates (n = 4986), 2968 studies remained. After screening the titles and abstracts, 69 full-text articles were left. A further 54 studies had to be excluded as they did not present data on injuries, included non-football players, or were neither cluster RCTs nor RCTs. Finally, 15 articles were included in the meta-analysis (Fig. 1).

3.2 Demographic and Study Characteristics

Eight trials stemmed from Europe [4, 9–11, 13, 32–34]. Two trials were conducted in the USA [8, 35]. One trial was

conducted in one each of the following countries: Canada [12], Australia [36], Rwanda [37], Nigeria [38], and Iran [39]. The overall number of participants was 22,177 including both sexes. Participants were registered football players in one of the following age groups: children (7-14 years), youth (12–19 years), senior, and veteran (> 32 years). The number of participants ranged from 265 [4] to 4564 participants [9]. A total of 5080 injuries and 1,587,327 h of exposure were included. The study period lasted between 12 weeks [8] and 9 months [4, 13, 39, 33]. All interventions were applied at least twice a week in the training sessions. The control groups performed their usual warm-up exercises and/or training routines. One study required an additional home-based stretching program [12]. Nine studies used a FIFA[®] warm-up program of the FIFA[®] 11, the FIFA[®] 11+, or the 11 + Kids [4, 10, 11, 13, 33, 35, 37-39]. Two studies used Neuromuscular Training programs [12, 32], and one study each used the Neuromuscular Control Program [36], the Knäkontroll program [9], the Prevention Injury and Enhance Performance program [8], and the Bounding Exercise Program [34] (Table 2).

3.3 Risk of Bias

Seven (46%) studies had a high risk of bias in two or more domains. The domain "other bias" was the most frequent cause for a high risk of bias within the studies (46%), with seven studies neither reporting an intention-to-treat analysis nor an adjustment for clustering (Fig. 1 and Table 1 of the Electronic Supplementary Material [ESM]).



Fig. 1 Flow chart of the included studies. RCTs randomized controlled trials

3.4 Meta-Analysis Results

3.4.1 Overall, Body Region, Contact, and Non-Contact-Related Injuries

For the primary outcome analysis, i.e., the overall injury risk, the pooled results showed a point estimate (RR) of 0.71 (95% CI 0.59–0.85; 95% PI 0.38–1.32; $I^2 = 80.5\%$; $\tau^2 = 0.067$; p < 0.001). The width of the 95% PI suggests that the effect in future similar studies lies between 0.38 and 1.32 (Fig. 2). In practical terms, the effect may vary from being very protective to an increased risk of injury. The level of evidence was rated as very low (downgraded one level because of a risk of bias, one level because of inconsistency, and one level because of publication bias) (Table 1).

Regarding the secondary outcome analyses, i.e., the body region-specific injury risk (Fig. 2), the point estimate (RR) for the lower limb injuries was 0.82 (95% CI 0.71–0.94; 95% PI 0.58–1.15; $I^2 = 45.3\%$; $\tau^2 = 0.016$; p = 0.067) with moderate-level evidence (downgraded one level because of a risk of bias). For knee injuries, the RR was 0.69 (95% CI 0.52–0.90; 95% PI 0.31–1.50) with low-level evidence

(downgraded one level because of a risk of bias and one level because of inconsistency). For hip/groin injuries, the RR was 0.56 (95% CI 0.30–1.05; 95% PI 0.00–102.92) with low-level evidence (downgraded one level because of a risk of bias and one level because of imprecision). For hamstring injuries, the RR was 0.83 (95% CI 0.50–1.37) with lowlevel evidence (downgraded one level because of a risk of bias and one level because of imprecision). With regard to ankle injuries, the RR was 0.73 (95% CI 0.55–0.96; 95% PI 0.36–1.46) with moderate-level evidence (downgraded one level because of a risk of bias). For each calculation, the 95% PI was wider in comparison to the 95% CI.

The pooled results for non-contact injuries showed a point estimate (RR) of 0.78 (95% CI 0.55–1.10; 95% PI 0.25–2.47; $I^2 = 67.3\%$; $\tau^2 = 0.100$; p = 0.016), with evidence rated as low level (downgraded one level because of a risk of bias and one level because of inconsistency). Additionally, the point estimate (RR) for contact injuries was 0.70 (95% CI 0.56–0.88; 95% PI 0.40–1.24 $I^2 = 29.2\%$; $\tau^2 = 0.018$; p = 0.227), with moderate-level evidence (downgraded one level because of a risk of bias). The width of the 95% PI suggested that the effect may vary from being very protective

 Table 2
 Summary of included multi-component randomized controlled trials investigating the effect of injury prevention programs

Study	Intervention program	Population (age)	Follow-up	Outcome	Number of ana- lysed (players)	Exposure time (h)	Number of injuries
Emery et al. 2010 [12]	Neuromuscular training program	Male and female youth (13– 18 years)	20 weeks	Overall injuries	IG: 380 CG: 364	IG: 24 051 CG: 24 597	IG: 50 CG: 79
Finch et al. 2016 [36]	Neuromuscular control program	Male senior (18–30 years)	28 weeks	Overall injuries	IG: 679 CG: 885	IG: 12 790 ^a CG: 15 537 ^a	IG: 335 CG: 438
Gilchrist et al. 2008 [8]	PEP	Female senior, (19.88 years) ^b	12 weeks	Knee injuries	IG: 583 CG: 852	IG: 35 220 CG: 52 919	IG: 40 ^c CG: 58 ^c
Hammes et al. 2015 [4]	FIFA [®] 11+	Male veteran $(\geq 32 \text{ years})$	9 months	Overall injuries	IG: 146 CG:119	IG: 4 172 CG: 2 937	IG: 51 CG: 37
Hilska et al. 2021 [32]	Neuromuscular training	Male and female children (9–14 years)	20 weeks	Lower limb inju- ries	IG: 673 CG: 730	IG: 71 109 CG: 63 404	IG: 310 ^d CG: 346 ^d
Nuhu et al. 2021 [37]	FIFA [®] 11+	Male senior (IG: 19.9 years) (CG: 19.7 years))	7 months	Overall injuries	IG: 309 CG: 317	IG: 65 333 CG: 63 389	IG: 168 CG: 252
Owoeye et al. 2014 [38]	FIFA [®] 11+	Male youth (14–19 years)	6 months	Overall injuries	IG: 212 CG: 204	IG: 51 017 CG: 61 045	IG: 36 CG: 94
Rossler et al. 2018 [13]	11 + Kids	Male and female children (7–13 years)	9 months	Overall injuries	IG: 2066 CG: 1829	IG: 140 716 CG: 152 033	IG: 139 CG: 235
Silvers-Granell et al. 2017 [35]	FIFA [®] 11+	Male senior (18–25 years)	5 months	Overall injuries	IG: 675 CG: 850	IG: 35 226 CG: 44 212	IG: 285 CG: 665
Soligard et al. 2008 [11]	FIFA [®] 11+	Female youth (13–17 years)	8 months	Overall injuries	IG: 1055 CG: 837	IG: 49 899 CG: 45 428	IG: 161 CG: 215
Steffen et al. 2008 [10]	FIFA [®] program 11	Female youth (13–17 years)	8 months	Overall injuries	IG: 1073 CG: 947	IG: 66 423 CG: 65 725	IG: 242 CG: 241
Walden et al. 2012 [9]	Knakontrol	Female youth (12–17 years)	7 months	ACL injuries	IG: 2479 CG: 2085	IG: 149 214 CG: 129 084	IG: 7 ^e CG: 14 ^e
Zarei et al. 2020 [39]	11 + kids	Male children (7–14 years)	9 months	Overall injuries	IG: 443 CG: 519	IG: 31 934 CG: 32 113	IG: 30 CG: 60
Van de Beijs- terveldt et al. 2012 [33]	FIFA [®] program 11	Male senior (18–40 years)	9 months	Overall injuries	IG: 233 CG: 233	IG: 21 605 CG: 22 647	IG: 207 CG: 220
Van de Hoef et al. 2019 [34]	BEP	Male senior (18–45 years)	39 weeks	Hamstring inju- ries	IG: 229 CG: 171	IG: 31 831 CG: 21 717	IG: 35 ^f CG: 30 ^f

ACL anterior cruciate ligament, BEP bounding exercise program, CG control group, IG intervention group, N/A Not applicable, PEP Prevent injury and Enhance Performance

^aMatch exposure only was reported

^bAverage age only was reported

^cKnee injuries

^dLower limb injuries

^eACL injuries

^fHamstring injuries

to an increased risk of injury for both outcomes, i.e., noncontact injuries (95% PI 0.55-1.10) and contact injuries (95% PI 0.40-1.24) (Fig. 3).

3.4.2 Subgroup Analysis According to Sex

Regarding a distinction between male and female individuals, the point estimate (RR) for the overall number of injuries

Injury Prevention in Football

Study	Year	Number of Injuries	Intervention	Control		% RR (95% CI) Weight
Overall injuries Soligard et al. [11] Steffen et al. [10] Emery et al. [12] van Beijsterveldt et al. [33] Owoeye et al. [38] Hammes et al. [4] Finch et al. [36] Silvers-Granelli et al. [35] Rossler et al. [13] Zarei et al. [39] Nuhu et al. [37] Subgroup, DL (1 ² = 80.5%, x ² = 0.0	2008 2008 2010 2012 2014 2015 2016 2017 2018 2020 2020 2020	376 483 129 437 139 88 773 950 374 90 420	161 242 50 207 36 51 335 285 139 30 168	215 241 79 220 94 37 438 665 235 60 252		$\begin{array}{c} 0.68 \ (0.56, 0.83) \ 11.32 \\ 1.00 \ (0.82, 1.22) \ 11.32 \\ 0.62 \ (0.39, 0.99) \ 7.13 \\ 0.97 \ (0.81, 1.16) \ 11.68 \\ 0.59 \ (0.36, 0.97) \ 6.70 \\ 0.91 \ (0.53, 1.57) \ 6.10 \\ 0.92 \ (0.68, 1.24) \ 9.79 \\ 0.54 \ (0.48, 0.61) \ 12.42 \\ 0.54 \ (0.48, 0.61) \ 7.03 \\ 0.50 \ (0.28, 0.89) \ 5.71 \\ 0.60 \ (0.47, 0.76) \ 10.81 \end{array}$
with estimated 95% predictive inte	rval				_ → +	0.71 (0.59, 0.85)100.00 (0.38, 1.32)
Ankle injuries Soligard et al. [11] Steffen et al. [10] Emery et al. [12] Owceye et al. [38] Rossler et al. [38] Nuhu et al. [37] Subgroup, DL (12 41.4%, τ ² = 0.0	2008 2008 2010 2014 2020 2020 2021 553, p = 0.115)	103 153 41 40 70 25 76	51 79 14 10 26 9 27	52 74 27 30 44 16 49		0.89 (0.61, 1.30) 22.26 1.10 (0.80, 1.51) 25.73 0.50 (0.24, 1.04) 10.51 0.53 (0.22, 1.30) 7.67 0.52 (0.22, 1.22) 8.41 0.57 (0.20, 1.65) 5.86 0.57 (0.37, 0.89) 19.55
with estimated 95% predictive inte Hamstring, injuries	rval				~	(0.36, 1.46)
Soligard et al. [11] van de Hoef et al. [34]	2008 2019	13 57	5 31	8 26		0.57 (0.18, 1.77) 19.26 0.91 (0.52, 1.58) 80.74
Subgroup, DL ($I^2 = 0.0\%$, $\tau^2 = 0.0\%$	0 , p = 0.470)				\diamond	0.83 (0.50, 1.37)100.00
Soligard et al. [11] Steffen et al. [10] Rossler et al. [13]	2008 2008 2018	19 20 20	10 6 4	9 14 16	#	1.01 (0.41, 2.49) 36.83 0.40 (0.17, 0.94) 40.10 0.40 (0.12, 1.34) 23.07
with estimated 95% predictive inte	erval				$\leftarrow \rightarrow$	0.56 (0.30, 1.05)100.00 (0.00, 102.92)
Knee injuries Gilchrist et al. [8] Soligard et al. [11] Steffen et al. [10] Ernery et al. [12] Waldén et al. [90 Owoeye et al. [38] Hammes et al. [4] Finch et al. [36] Silvers-Granelli et al. [35] Rossier et al. [13] Nuhu et al. [37]	2008 2008 2010 2012 2014 2015 2016 2017 2018 2021	98 93 67 11 21 33 10 92 136 83 83	40 35 37 3 7 12 6 32 34 29 38	58 58 30 8 14 21 4 60 102 54 43		1.04 (0.70, 1.56) 14.00 0.55 (0.36, 0.84) 13.57 1.20 (0.76, 1.90) 12.87 0.38 (0.08, 1.78) 2.76 0.38 (0.15, 0.86) 6.71 0.93 (0.38, 2.29) 6.38 1.06 (0.21, 5.46) 2.48 0.50 (0.24, 1.05) 8.22 0.42 (0.26, 0.68) 12.37 0.47 (0.20, 1.15) 6.54 0.91 (0.61, 1.35) 14.10
Subgroup, DL ($I^2 = 52.4\%$, $\tau^2 = 0.0$ with estimated 95% predictive inte	199, p = 0.021) erval				-~	0.69 (0.52, 0.90)100.00
Lower limb injuries Soligard et al. [11] Steffen et al. [10] Emery et al. [12] Owoeye et al. [38] Hammes et al. [4] Finch et al. [36] Zarei et al. [39] Hilska et al. [39] Nuhu et al. [37] Subgroup, DL. $(l^2 = 45.3\%, \tau^2 = 0.0)$ with estimated 95% predictive inte	2008 2010 2014 2015 2016 2020 2021 2021 2021 2021 2021 2021	264 354 102 37 379 78 656 260	121 181 42 26 43 151 24 310 116	143 173 60 76 30 228 54 346 144		0.71 (0.49, 1.03) 9.46 1.00 (0.78, 1.27) 15.66 0.68 (0.42, 1.11) 6.34 0.52 (0.29, 0.92) 4.85 1.01 (0.55, 1.86) 4.35 0.78 (0.56, 1.06) 11.14 0.45 (0.24, 0.83) 4.30 0.82 (0.64, 1.05) 15.66 0.94 (0.87, 1.01) 28.25 0.82 (0.71, 0.94)100.00 (0.58, 1.15)
					I I .1 1 2	
					Favours intervention Favours control	

Fig.2 Analysis of multi-component exercise-based injury prevention programs' effect on the overall and region-specific injury risk compared with control groups. $I^2 I$ square, p p value, RR risk ratio, τ^2 tau square

in male football players was 0.70 (95% CI 0.55–0.90; $I^2 = 83.5\%$; $\tau^2 = 0.082$; p < 001). In female football players, the point estimate (RR) was 0.82 (95% CI 0.57–1.20; $I^2 = 68.9\%$; $\tau^2 = 0.064$; p = 0.008) (Fig. 4 of the ESM).

3.4.3 Subgroup Analysis According to Age Group

The point estimate (RR) for the overall number of injuries in children was 0.52 (95% CI 0.36–0.76; $I^2 = 0.0\%$; $\tau^2 < 0.001$; p = 0.841), in youth, the RR was 0.74 (95% CI 0.56–0.97; $I^2 = 68.9\%$; $\tau^2 = 0.048$; p = 0.022), in seniors, the RR was 0.73 (95% CI 0.53–1.01; $I^2 = 91.1\%$; $\tau^2 = 0.098$; p < 0.001),

and, in veterans, the RR was 0.91 (95% CI 0.53–1.57) (Fig. 4 of the ESM).

4 Discussion

4.1 Principal Findings

This systematic review and meta-analysis included 15 RCTs that assessed the effect of injury prevention programs on the overall and body region-specific injury risk in football players. Based on calculated PIs, their efficacy remains uncertain



Fig. 3 Analysis of multi-component exercise-based injury prevention programs' effect on the overall non-contact (a) and contact (b) injury risk compared with control groups. I^2 I square, p p value, RR risk ratio, τ^2 tau square

and inconclusive regarding all primary and secondary outcomes. In addition, the majority of the results are based on low-quality evidence.

4.1.1 Comparison with Existing Literature on Injury Risk Reduction

Riley et al. [40] suggested that if a random-effects approach is used, the pooled result must be interpreted as the average intervention effect across studies, rather than the common effect. Previous meta-analyses have not reported PIs, which means, an appropriate comparison is not possible. Therefore, we can only compare our point estimates with those reported in the literature. In contrast with the currently available evidence [14–16, 41], our study included footballers of all age groups and skill levels (amateur and professional). The point estimate (RR) of 0.71 (95% CI 0.59-0.85) in the current analysis is at the lower end of those reported in previous systematic reviews, which reported an incidence rate ratio (IRR) of 0.73 (95% CI 0.59–0.91) [41], IRR of 0.75 (95% CI 0.57–0.98) [14], IRR of 0.77 (95% CI 0.64-0.91) [15], and IRR of 0.77 (95% CI 0.61–0.97) [16]. This was to be expected as we also included interventions in children, which showed a substantially higher injury reduction of 48% [13] and 50% [39] compared with older players. This effect was somewhat counterbalanced by the reduced effect of the programs among veterans, which was only 9%. However, the relative weight of the studies with children was higher (higher in the number of studies and participants). A previous meta-analysis [14] investigated the effect of the FIFA® exercise-based injury prevention programs on specific body regions. The observed efficacy on hamstring (RR 0.83 vs IRR 0.40), knee (RR 0.69 vs IRR 0.52), and ankle injuries (RR 0.73 vs IRR 0.68) was lower in our study, but comparable for hip/groin injuries (RR 0.56 vs IRR 0.59). A likely explanation for the differing results between the reviews is that we included a higher number of studies that examined different types of programs in the analysis. An additional explanation could be the inclusion of studies with children because injury patterns vary with age [42]. The most obvious difference from other studies was regarding hamstring injuries. The results may be expected as we did not include trials investigating the Nordic Hamstring as a single component exercise, which has been shown to be very effective for preventing hamstring injuries [43]. Moreover, in comparison to Thorborg et al. [14], we included the Bounding Exercise Program [34], which showed very little effect in reducing these injuries.

4.1.2 Effectiveness of Injury Prevention Programs on Contact Versus Non-contact Injuries

For the first time, this study investigated the effect of multicomponent exercise-based injury prevention programs not only on non-contact injuries but also on contact-related injuries. The point estimate (RR) for contact injuries was 0.70 (95% CI 0.56–0.88). Surprisingly, the estimated risk reduction was higher than for non-contact injuries for which the vast majority of programs are designed. Most programs include strength exercises that mostly focus on core stability. Furthermore, plyometrics (hopping, jumping, and landing) are often part of the programs. They have the potential to improve lower leg strength, functional leg stability, and balance, thus improving the ability to absorb external forces, for example, induced by contact. The 11 + Kids [13] program also includes one exercise specifically on correct falling techniques. The point estimate (RR) for non-contact injuries in the current study was 0.78, in line with a previous study that reported a RR of 0.77 [16].

4.1.3 Effectiveness of Injury Prevention Programs Across Sexes and Age Groups

The subgroup analysis showed a point estimate (RR) of 0.70 in male football players. These results mimic the data of the Al Attar et al. study [15]. However, the estimated effect is slightly lower than data reported by Lemes et al. [16] showing a point estimate (RR) of 0.68.

Regarding female individuals, the pooled results showed a point estimate (RR) of 0.82. This result falls within the range of results reported by studies with similar inclusion criteria [15, 16]. However, the meta-analysis with the largest estimated effect [41] included RCTs that used various injury prevention strategies. In addition to physical exercises, they included studies that used braces and education as a method for prevention. Furthermore, they included studies with participants of varying backgrounds and sports (i.e., middle and high school non-footballer athletes). These dissimilarities might have caused these considerable differences. In contrast, small differences compared with other reviews [15, 16] may reflect the diversity of interventions, i.e., the inclusion of single-component exercise-based injury prevention programs.

The subgroup analysis for age groups showed a point estimate (RR) of 0.52 in children, a RR of 0.74 in youth, 0.73 in seniors, and 0.91 in veteran football players. The point estimate in youth and seniors is homogeneous with the current available evidence [14, 41]. The low point estimate found in children may be expected by the fact that there is rarely any prior use of preventative measures at all; therefore, using the program is likely to evoke the biggest benefit. Only one trial [4] assessed the effects of injury prevention programs in veteran football players. The comparably small effect in this population is likely owing to the infrequent application of the program (only once a week) as well as relatively low compliance.

4.2 Factors to Take into Account When Assessing PIs

In the current analysis, we calculated the PIs for the main investigated outcomes. Prediction intervals were wider in comparison to confidence intervals. Based on this evidence, there is a lack of compelling data to affirm the certainty of preventive effects from multi-component exercise-based injury prevention programs. However, for our meta-analysis, we have to take into account that the use of PIs has its shortcomings. IntHout et al. [19] mentioned that they show a wider range compared with CIs when there is any heterogeneity. Our main outcome provided an $I^2 = 80.5\%$, which should be interpreted as high heterogeneity according to the Cochrane Handbook for Systematic Reviews of Interventions [27]. In addition, Riley et al. [40] stated that a PI will be most appropriate when the studies included in the meta-analysis have a low risk of bias. However, the majority of studies in our analysis had a high risk of bias. Therefore, these shortcomings would have affected the use of PIs in our meta-analysis.

4.3 Strengths and Limitations

To the best of our knowledge, this review is the first to analyze the efficacy of multi-component exercise-based injury prevention programs among footballers of all age groups. One strength of this systematic review is that it included multiple analyses. It investigated the risk reduction for the overall number of injuries as well as of body region-specific, contact, and non-contact injuries. Subgroup analyses for age and sex were also performed. Additionally, the PIs for the main outcomes were calculated. A further strength is the large number of participants (22,177), injuries (5080), and exposure hours (1,587,327 h) included in comparison with other reviews [14–16]. Furthermore, we followed best practice by including only randomized trials and cluster-RCTs, using a risk of bias assessment and grading the quality of evidence.

However, this review also has some limitations, mainly that > 50% of the reported effects were based on studies with a very low or low level of evidence. The main outcome variable provided high heterogeneity among the studies $(I^2 = 80.5\%)$. The lack of information about compliance with the prevention program in many studies is another limitation of this review. Furthermore, there was missing information on content and compliance with the usual warm-ups/training routines of the control groups. Another limitation is the high risk of bias, especially from the "other bias" domain, with seven studies failing to report the use of an intention-to-treat analysis and of an adjustment for clustering. Finally, two deviations (lack of a compliance analysis and the modification of literature databases) from the original study protocol have to be mentioned as limitations of this review.

4.4 Differences Between the Protocol and Review

Owing to the lack of respective information provided in the studies, a compliance analysis was impossible. We contacted the corresponding authors to provide us with these data, but within the set time of 2 weeks, we only received information on one of the studies. Our planned bibliographic databases for literature identification were modified during the study implementation. Because of the lack of access, we did not search in EMBASE and SPORT-Discus. However, we additionally searched in the originally unplanned database Scopus. In addition, to empower the review, although it was not registered in the protocol, we assessed the quality of evidence using the GRADE approach and calculated the PIs for the main outcomes.

4.5 Recommendations for Future Studies

Based on the data obtained, we recommend future highquality trials to investigate the efficacy of multi-component exercise-based injury prevention programs. In upcoming studies, data on compliance and the content of the training of the control groups should be included. Adjustment for clustering and more extensive reporting of outcomes should be emphasized. In addition, it appears important to create new injury prevention programs that reflect the development and changes in football training. This should include increasing their attractiveness to promote compliance (also outside of study settings), which appears crucial to reduce injury risk. Currently, a large number of different exercises are included because it is unknown which exercises (or which combination of them) are most effective in general or in relation to specific injuries. Tailoring the exercises would potentially mean fewer injuries and more efficiency.

5 Conclusions

This meta-analysis indicated that evidence for the meaningful effects of multi-component exercise-based injury prevention programs in football remains inconclusive at best. This statement is based on PIs that were wider than the frequently employed CIs, with a range from very protective effects to an increased injury risk. In addition, the quality of evidence is a major issue in existing studies. These findings call for future high-quality trials to provide more reliable evidence regarding the efficacy of injury prevention programs in football. Supplementary Information The online version contains supplementary material available at https://doi.org/10.1007/s40279-022-01797-7.

Declarations

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Authors' contributions RO, IH, and KadF conceived and designed the study. RO, IH, RM, AL, AB, DK, and BSH conducted the search, study selection, data extraction, and quality appraisal. IH analyzed the data. RO, IH, and KadF contributed to the interpretation of the data. RO drafted the manuscript with input from KadF and TM. All authors have read and approved the final manuscript.

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References

- Aus der Fünten K, Faude O, Lensch J, Meyer T. Injury characteristics in the German professional male soccer leagues after a shortened winter break. J Athl Train. 2014;49(6):786–93. https:// doi.org/10.4085/1062-6050-49.3.51.
- van Beijsterveldt AM, Stubbe JH, Schmikli SL, van de Port IG, Backx FJ. Differences in injury risk and characteristics between Dutch amateur and professional soccer players. J Sci Med Sport. 2015;18(2):145–9. https://doi.org/10.1016/j.jsams.2014.02.004.
- Szymski D, Krutsch V, Achenbach L, Gerling S, Pfeifer C, Alt V, et al. Epidemiological analysis of injury occurrence and current prevention strategies on international amateur football level during the UEFA Regions Cup 2019. Arch Orthop

Trauma Surg. 2022;142(2):271-80. https://doi.org/10.1007/ s00402-021-03861-9.

- Hammes D, Aus der Fünten K, Kaiser S, Frisen E, Bizzini M, Meyer T. Injury prevention in male veteran football players: a randomised controlled trial using "FIFA 11+". J Sports Sci. 2015;33(9):873–81. https://doi.org/10.1080/02640414.2014. 975736.
- Faude O, Rößler R, Junge A. Football injuries in children and adolescent players: are there clues for prevention? Sports Med. 2013;43(9):819–37. https://doi.org/10.1007/s40279-013-0061-x.
- Ekstrand J, Hägglund M, Waldén M. Injury incidence and injury patterns in professional football: the UEFA injury study. Br J Sports Med. 2011;45(7):553–8. https://doi.org/10.1136/bjsm. 2009.060582.
- Materne O, Chamari K, Farooq A, Weir A, Hölmich P, Bahr R, et al. Injury incidence and burden in a youth elite football academy: a four-season prospective study of 551 players aged from under 9 to under 19 years. Br J Sports Med. 2021;55(9):493–500. https://doi.org/10.1136/bjsports-2020-102859.
- Gilchrist J, Mandelbaum BR, Melancon H, Ryan GW, Silvers HJ, Griffin LY, et al. A randomized controlled trial to prevent noncontact anterior cruciate ligament injury in female collegiate soccer players. Am J Sports Med. 2008;36(8):1476–83. https://doi.org/ 10.1177/0363546508318188
- Waldén M, Atroshi I, Magnusson H, Wagner P, Hägglund M. Prevention of acute knee injuries in adolescent female football players: cluster randomised controlled trial. BMJ. 2012;344: e3042. https://doi.org/10.1136/bmj.e3042.
- Steffen K, Myklebust G, Olsen OE, Holme I, Bahr R. Preventing injuries in female youth football—a cluster-randomized controlled trial. Scand J Med Sci Sports. 2008;18(5):605–14. https://doi.org/ 10.1111/j.1600-0838.2007.00703.x.
- Soligard T, Myklebust G, Steffen K, Holme I, Silvers H, Bizzini M, et al. Comprehensive warm-up programme to prevent injuries in young female footballers: cluster randomised controlled trial. BMJ. 2008;9(337): a2469. https://doi.org/10.1136/bmj.a2469.
- Emery CA, Meeuwisse WH. The effectiveness of a neuromuscular prevention strategy to reduce injuries in youth soccer: a clusterrandomised controlled trial. Br J Sports Med. 2010;44(8):555–62. https://doi.org/10.1136/bjsm.2010.074377.
- Rössler R, Junge A, Bizzini M, Verhagen E, Chomiak J, Aus der Fünten K, et al. A multinational cluster randomised controlled trial to assess the efficacy of "11+ Kids": a warm-up programme to prevent injuries in children's football. Sports Med. 2018;48(6):1493– 504. https://doi.org/10.1007/s40279-017-0834-8.
- 14. Thorborg K, Krommes KK, Esteve E, Clausen MB, Bartels EM, Rathleff MS. Effect of specific exercise-based football injury prevention programmes on the overall injury rate in football: a systematic review and meta-analysis of the FIFA 11 and 11 + programmes. Br J Sports Med. 2017;51(7):562–71. https://doi. org/10.1136/bjsports-2016-097066.
- Al Attar WS, Soomro N, Pappas E, Sinclair PJ, Sanders RH. How effective are F-Marc Injury Prevention programs for soccer players? A systematic review and meta-analysis. Sports Med. 2015;46(2):205–17. https://doi.org/10.1007/s40279-015-0404-x.
- Lemes IR, Pinto RZ, Lage VN, Roch BAB, Verhagen E, Bolling C, et al. Do exercise-based prevention programmes reduce noncontact musculoskeletal injuries in football (soccer)? A systematic review and meta-analysis with 13 355 athletes and more than 1 million exposure hours. Br J Sports Med. 2021;55(20):1170–8. https://doi.org/10.1136/bjsports-2020-103683.
- López-Valenciano A, Ruiz-Pérez I, Garcia-Gómez A, Vera-Garcia FJ, De Ste CM, Myer GD, et al. Epidemiology of injuries in professional football: a systematic review and meta-analysis. Br J Sports Med. 2020;54(12):711–8. https://doi.org/10.1136/bjspo rts-2018-099577.

- Impellizzeri FM, McCall A, van Smeden M. Why methods matter in a meta-analysis: a reappraisal showed inconclusive injury preventive effect of Nordic hamstring exercise. J Clin Epidemiol. 2021;140:111–24. https://doi.org/10.1016/j.jclinepi.2021.09.007.
- IntHout J, Ioannidis JP, Rovers MM, Goeman JJ. Plea for routinely presenting prediction intervals in meta-analysis. BMJ Open. 2016;6(7): e010247. https://doi.org/10.1136/bmjop en-2015-010247.
- Page MJ, McKenzie JE, Bossuyt PM, Boutron I, Hoffmann TC, Mulrow CD, et al. The PRISMA 2020 statement: an updated guideline for reporting systematic reviews. BMJ. 2021;372: n71. https://doi.org/10.1136/bmj.n71.
- Higgins JP, Altman DG, Gøtzsche PC, Jüni P, Moher D, Oxman AD, Cochrane Bias Methods Group; Cochrane Statistical Methods Group, et al. The Cochrane Collaboration's tool for assessing risk of bias in randomised trials. BMJ. 2011;343:d5928. https://doi. org/10.1136/bmj.d5928.
- Atkins D, Best D, Briss PA, Eccles M, Falck-Ytter Y, Flottorp S, GRADE Working Group, et al. Grading quality of evidence and strength of recommendations. BMJ. 2004;328(7454):1490. https:// doi.org/10.1136/bmj.328.7454.1490.
- Meader N, King K, Llewellyn A, Norman G, Brown J, Rodgers M, et al. A checklist designed to aid consistency and reproducibility of GRADE assessments: development and pilot validation. Syst Rev. 2014;3(1):1–9. https://doi.org/10.1186/ 2046-4053-3-82.
- Guyatt GH, Oxman AD, Kunz R, Vist GE, Falck-Ytter Y, Schünemann HJ, GRADE Working Group. What is "quality of evidence" and why is it important to clinicians? BMJ. 2008;336(7651):995–8. https://doi.org/10.1136/bmj.39490. 551019.BE.
- 25. Grant RL. Converting an odds ratio to a range of plausible relative risks for better communication of research findings. BMJ. 2014;348: f7450. https://doi.org/10.1136/bmj.f7450.
- Shor E, Roelfs D, Vang ZM. The "Hispanic mortality paradox" revisited: meta-analysis and meta-regression of life-course differentials in Latin American and Caribbean immigrants' mortality. Soc Sci Med. 2017;186:20–33. https://doi.org/10.1016/j. socscimed.2017.05.049.
- Higgins JPT, Green S. Cochrane handbook for systematic reviews of interventions. Version 5.1.0 [updated March 2011]. The Cochrane Collaboration; 2011.
- Fawzi WW, Chalmers TC, Herrera MG, Mosteller F. Vitamin A supplementation and child mortality: a meta-analysis. JAMA. 1993;269(7):898–903. https://doi.org/10.1001/jama.1993.03500 070078033.
- 29. Egger M, Davey-Smith G, Altman D, editors. Systematic reviews in health care: meta-analysis in context. New York: Wiley; 2008.
- Borenstein M, Hedges LV, Higgins JP, Rothstein HR. Introduction to meta-analysis. New York: Wiley; 2011.
- Egger M, Davey Smith G, Schneider M, Minder C. Bias in meta-analysis detected by a simple, graphical test. BMJ. 1997;315(7109):629–34. https://doi.org/10.1136/bmj.315.7109. 629.
- 32. Hilska M, Leppänen M, Vasankari T, Aaltonen S, Kannus P, Parkkari J, et al. Neuromuscular training warm-up prevents acute noncontact lower extremity injuries in children's soccer: a cluster randomized controlled trial. Orthop J Sports Med. 2021;9(4):23259671211005770. https://doi.org/10.1177/23259 671211005769.
- 33. van Beijsterveldt AM, van de Port IG, Krist MR, Schmikli SL, Stubbe JH, Frederiks JE, et al. Effectiveness of an injury prevention programme for adult male amateur soccer players: a clusterrandomised controlled trial. Br J Sports Med. 2012;46(16):1114– 8. https://doi.org/10.1136/bjsports-2012-091277.

- 34. van de Hoef PA, Brink MS, Huisstede BMA, van Smeden M, de Vries N, Goedhart EA, et al. Does a bounding exercise program prevent hamstring injuries in adult male soccer players? A cluster-RCT. Scand J Med Sci Sports. 2019;29(4):515–23. https://doi.org/ 10.1111/sms.13353.
- Silvers-Granelli HJ, Bizzini M, Arundale A, Mandelbaum BR, Snyder-Mackler L. Does the FIFA 11 + injury prevention program reduce the incidence of ACL injury in male soccer players? Clin Orthop Relat Res. 2017;475(10):2447–55. https://doi.org/10. 1007/s11999-017-5342-5.
- 36. Finch CF, Twomey DM, Fortington LV, Doyle TL, Elliott BC, Akram M, et al. Preventing Australian football injuries with a targeted neuromuscular control exercise programme: comparative injury rates from a training intervention delivered in a clustered randomised controlled trial. Inj Prev. 2016;22(2):123–8. https:// doi.org/10.1136/injuryprev-2015-041667.
- Nuhu A, Jelsma J, Dunleavy K, Burgess T. Effect of the FIFA 11 + soccer specific warm up programme on the incidence of injuries: a cluster-randomised controlled trial. PLoS ONE. 2021;16(5): e0251839. https://doi.org/10.1371/journal.pone.0251839.
- Owoeye OB, Akinbo SR, Tella BA, Olawale OA. Efficacy of the FIFA 11 + warm-up programme in male youth football: a cluster randomised controlled trial. J Sports Sci Med. 2014;13(2):321–8.
- Zarei M, Abbasi H, Namazi P, Asgari M, Rommers N, Rössler R. The 11+ Kids warm-up programme to prevent injuries in young

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Iranian male high-level football (soccer) players: a cluster-randomised controlled trial. J Sci Med Sport. 2020;23(5):469–74. https://doi.org/10.1016/j.jsams.2019.12.001.

- 40. Riley RD, Higgins JP, Deeks JJ. Interpretation of random effects meta-analyses. BMJ. 2011;342: d549. https://doi.org/10. 1136/bmj.d549.
- 41. Crossley KM, Patterson BE, Culvenor AG, Bruder AM, Mosler AB, Mentiplay BF. Making football safer for women: a systematic review and meta-analysis of injury prevention programmes in 11 773 female football (soccer) players. Br J Sports Med. 2020;54(18):1089–98. https://doi.org/10.1136/bjspo rts-2019-101587.
- 42. Wik EH, Lolli L, Chamari K, Materne O, Di Salvo V, Gregson W, et al. Injury patterns differ with age in male youth football: a four-season prospective study of 1111 time-loss injuries in an elite national academy. Br J Sports Med. 2021;55(14):794–800. https://doi.org/10.1136/bjsports-2020-103430.
- 43. van der Horst N, Smits DW, Petersen J, Goedhart EA, Backx FJ. The preventive effect of the Nordic hamstring exercise on hamstring injuries in amateur soccer players: a randomized controlled trial. Am J Sports Med. 2015;43(6):1316–23. https://doi.org/10. 1177/0363546515574057.

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